

# Consistent On-Line Off-Policy Evaluation

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## Abstract

The problem of on-line off-policy evaluation (OPE) has been actively studied in the last decade due to its importance both as a stand-alone problem and as a sub-module in a policy improvement scheme. However, most Temporal Difference (TD) based solutions ignore the discrepancy between the stationary distribution of the behavior and target policies and its effect on the convergence limit when linear function approximation is applied. In this paper we propose Consistent Off-Policy TD (COP-TD) that addresses this issue directly and enables reducing this bias at some computational expense. We show that COP-TD( $\lambda, \beta$ ) can be designed to converge to the same value that would have been obtained by using on-policy TD( $\lambda$ ) with the target policy. Subsequently, the proposed scheme leads to a related and promising heuristic we call log-COP-TD( $\lambda, \beta$ ). Both algorithms have favorable empirical results to the current state of the art on-line OPE algorithms.

## 1. Introduction

In this paper we address the problem of assessing the performance of a complex strategy without trying it known as off-policy evaluation (OPE). OPE formulation is often considered in domains with limited sampling capability. For example, marketing and recommender systems (Theocharous and Hallak, 2013; Theocharous et al., 2015) directly relate policies to revenue. A more extreme example is drug administration, where there are only few patients in the testing population, and sub-optimal policies can have life threatening effects (Hochberg et al., 2016).

We consider the OPE problem in an on-line setup where each new sample is immediately used to update our current value estimate of some previously unseen policy. We propose a new algorithm called COP-TD( $\lambda, \beta$ ) for estimating the value of the target policy which is consistent with sufficient resources: it converges to the value on-policy learning with the target policy would have converged to. In addition, we introduce a related heuristic called Log-COP-TD( $\lambda, \beta$ ) and explain its motivation.

## 2. Notations and Background

We consider the standard discounted Markov Decision Process (MDP) formulation Bertsekas and Tsitsiklis (1996) with a single long trajectory. Let  $M = (\mathcal{S}, \mathcal{A}, \mathcal{P}, \mathcal{R}, \zeta, \gamma)$  be an MDP where  $\mathcal{S}$  is the finite state space and  $\mathcal{A}$  is the finite action space. The parameter  $\mathcal{P}$  sets the transition probabilities  $P(s'|s, a)$  given the previous state  $s \in \mathcal{S}$  and action  $a \in \mathcal{A}$ , where the first state is determined by the distribution  $\zeta$ . The parameter  $\mathcal{R}$  sets the reward distribution  $r(s, a)$  obtained by taking action  $a$  in state  $s$  and  $\gamma$  is the discount factor specifying the exponential reduction in reward with time.

The process advances as follows: A state  $s_0$  is sampled according to the distribution  $\zeta(s)$ . Then, each time step  $t$  starting from  $t = 0$  the agent chooses an action  $a_t$  according to some distribution

(defined as the policy)  $\mu(a|s_t)$  (the behavior policy), a reward  $r_t \doteq r(s_t, a_t)$  is accumulated by the agent, and the next state  $s_{t+1}$  is sampled using the transition probability  $P(s'|s_t, a_t)$ .

The expected discounted accumulated reward starting from a specific state and choosing an action by some policy  $\pi$  is called the value function, which is also known to hold the Bellman equation in a vector form:  $V^\pi(s) = \mathbb{E}_\pi [\sum_{t=0}^{\infty} \gamma^t r_t]$ ,  $T_\pi V \doteq R_\pi + \gamma P_\pi V$ , where  $[R_\pi]_s \doteq \mathbb{E}_\pi [r(s, \pi(s))]$  and  $[P_\pi]_{s,s'} \doteq \mathbb{E}_\pi [P(s'|s, \pi(s))]$  are the policy induced reward vector and transition probability matrix respectively;  $T_\pi$  is called the Bellman operator. The problem of estimating  $V^\pi(s)$  from samples is called policy evaluation. If the target policy  $\pi$  is different than the behavior policy  $\mu$  which generated the samples, the problem is called off-policy evaluation (OPE).

The TD( $\lambda$ ) (Sutton, 1988) algorithm is a standard solution to on-line on-policy evaluation: Each time step the TD error updates the current value function estimate, so eventually the process will converge to the true value function. The standard form of TD( $\lambda$ ) is given by:

$$\begin{aligned} R_{t,s_t}^{(n)} &= \sum_{i=0}^{n-1} \gamma^i r_{t+i} + \gamma^n \hat{V}_t(s_{t+n}), & R_{t,s_t}^\lambda &= (1 - \lambda) \sum_{n=0}^{\infty} \lambda^n R_{s_t}^{(n+1)}, \\ \hat{V}_{t+1}(s_t) &= \hat{V}_t(s_t) + \alpha_t \left( R_{t,s_t}^\lambda - \hat{V}_t(s_t) \right), \end{aligned} \quad (1)$$

where  $\alpha_t$  is the learning rate. The value  $R_{t,s_t}^{(n)}$  is an estimate of  $V(s_t)$ , looking forward  $n$  steps, and  $R_{t,s_t}^\lambda$  is an exponentially weighted average of all of these estimates going forward till infinity.

We denote by  $d_\mu(s)$  the stationary distribution over states induced by taking the policy  $\mu$  and mark  $D_\mu = \text{diag}(d_\mu)$ . Since we are concerned with the behavior at the infinite horizon, we assume  $\zeta(s) = d_\mu(s)$ . In addition, we assume the MDP is ergodic for the two specified policies  $\mu, \pi$  so  $\forall s \in \mathcal{S} : d_\mu(s) > 0, d_\pi(s) > 0$  and that the OPE problem is proper -  $\mu(a|s) > 0 \iff \pi(a|s) > 0$ .

When the state space is too large to hold  $V^\pi(s)$ , a linear function approximation scheme is used:  $V^\pi(s) \approx \theta_\pi^\top \phi(s)$ , where  $\theta$  is the optimized weight vector and  $\phi(s)$  is the feature vector of state  $s$  composed of  $k$  features; in this paper we assume the features are linearly independent. TD( $\lambda$ ) can be adjusted accordingly to find the fixed point of  $\Pi_{d_\pi} T_\pi^\lambda$  where  $\Pi_{d_\pi}$  is the projection to the subspace spanned by the features with respect to the  $d_\pi$ -weighted norm:

$$\begin{aligned} R_{t,s_t}^{(n)} &= \sum_{i=0}^{n-1} \gamma^i r_{t+i} + \gamma^n \theta_t^\top \phi(s_{t+n}), & R_{t,s_t}^\lambda &= (1 - \lambda) \sum_{n=0}^{\infty} \lambda^n R_{s_t}^{(n+1)} \\ \theta_{t+1} &= \theta_t + \alpha_t \left( R_{t,s_t}^\lambda - \theta_t^\top \phi(s_t) \right) \phi(s_t). \end{aligned} \quad (2)$$

Finally, we define OPE-related quantities:  $\rho_t \doteq \frac{\pi(a_t|s_t)}{\mu(a_t|s_t)}$ ,  $\Gamma_t^n \doteq \prod_{i=0}^{n-1} \rho_{t-1-i}$ ,  $\rho_d(s) \doteq \frac{d_\pi(s)}{d_\mu(s)}$ , we call  $\rho_d$  the stationary distribution ratio.

### 3. Previous Work

We can roughly categorize previous OPE algorithms to two main families: Gradient based methods that perform stochastic gradient descent on error terms they want to minimize (Sutton et al., 2009a,b; White and White, 2016), and Importance Sampling (IS) methods that correct the gains between on-policy and off-policy updates using the IS-ratios  $\rho_t$  (Precup et al., 2001; Sutton et al., 2015).

There are also a few algorithms that fall between the two, (van Hasselt et al., 2014; Mahmood and Sutton, 2015). Our work falls within the IS family.

A comparison of these algorithms in terms of convergence rate is available in (White and White, 2016; Geist and Scherrer, 2014). We focus on the limit point of the convergence. For most of the TD algorithms, the process was shown to converge almost surely to the fixed point of the projected Bellman operator  $\Pi_d T_\pi$  where  $d$  is some stationary distribution (usually  $d_\mu$ ), however the  $d$  in question was never<sup>1</sup>  $d_\pi$  as we would have obtained from running on-policy TD with the target policy. The algorithm achieving the closest result is ETD( $\lambda, \beta$ ) which replaced  $d$  with  $f = (I - \beta P_\pi^\top)^{-1} d_\mu$ , where  $\beta$  trades-off some of the process' variance with the bias in the limit point.

#### 4. COP-TD( $\lambda, \beta$ )

Most off-policy algorithms multiply the TD summand of TD( $\lambda$ ) with some value that depends on the history and the current state. For example, full IS-TD by Precup et al. (2001) examines the ratio between the trajectory's probabilities under both policies:  $\frac{P_\pi(s_0, a_0, s_1, \dots, s_t, a_t)}{P_\mu(s_0, a_0, s_1, \dots, s_t, a_t)} = \prod_{m=0}^t \rho_m = \Gamma_t^t \rho_t$ .

In problems with a long horizon, or these that start from the stationary distribution, we suggest using a time-invariant term which is the stationary distribution ratio  $\rho_d$  multiplied by the current  $\rho_t$ . This leads us to the following update equations:

$$\theta_{t+1} = \theta_t + \alpha_t \rho_d(s_t) \rho_t \left( r_t + \theta_t^\top (\gamma \phi(s_{t+1}) - \phi(s_t)) \right) \phi(s_t). \quad (3)$$

**Lemma 1** *If the step sizes  $\alpha_t$  satisfy  $\sum_{t=0}^{\infty} \alpha_t = \infty$ ,  $\sum_{t=0}^{\infty} \alpha_t^2 < \infty$  then the process described by Equation 3 converges almost surely to the fixed point of  $\Pi_\pi T_\pi V = V$ .*

In practice, we have no way of knowing  $\rho_d(s)$ , therefore we suggest estimating it using an additional stochastic approximation process. In order to do so, we note that the quantity  $\Gamma_t^n$  that is commonly used in IS-based algorithm satisfies the following, for any function on the state space  $u(s)$ :  $\mathbb{E}_\mu [\Gamma_t^n u(s_{t-n}) | s_t] = u^\top D_\mu P_\pi^n D_\mu^{-1} e_{s_t}$ , where  $e_{s_t}$  is the unit vector of state  $s_t$ . So, for an unbiased estimate of  $\rho_d$  denoted  $\hat{\rho}_d$  we can define and derive:  $\tilde{\Gamma}_t^n \doteq \hat{\rho}_d(s_{t-n}) \Gamma_t^n = \hat{\rho}_d(s_{t-n}) \prod_{i=0}^{n-1} \rho_{t-i-1} \Rightarrow \mathbb{E}_\mu [\tilde{\Gamma}_t^n | s_t] = \rho_d(s_t)$ .

For any state  $s_t$  there are  $t \rightarrow \infty$  such quantities  $\{\tilde{\Gamma}_t^n\}_{n=0}^t$ , where we propose to weight them similarly to TD( $\lambda$ ):  $\tilde{\Gamma}^\beta = (1 - \beta) \sum_{n=0}^{\infty} \beta^n \tilde{\Gamma}_t^{n+1}$ .

Note that  $\rho_d(s)$ , unlike  $V(s)$ , is restricted to a close set since its  $d_\mu$ -weighted linear combination is equal to 1 and all of its entries are non-negative; We denote this  $d_\mu$ -weighted simplex by  $\Delta_{d_\mu}$ , and  $\Pi_{\Delta_{d_\mu}}$  is the (non-linear) projection to this set with respect to the Euclidean norm ( $\Pi_{\Delta_{d_\mu}}$  can be calculated efficiently, Chen and Ye (2011)). Now, we can devise a TD algorithm which estimates  $\rho_d$  and use it to find  $\theta$ , which we call COP-TD(0,  $\beta$ ) (Consistent Off-Policy TD).

**Theorem 2** *If the step sizes satisfy  $\sum_t \alpha_t = \sum_t \alpha_t^d = \infty$ ,  $\sum_t (\alpha_t^2 + (\alpha_t^d)^2) < \infty$ ,  $\frac{\alpha_t}{\alpha_t^d} \rightarrow 0$ ,  $t \alpha_t^d \rightarrow 0$ , and  $\mathbb{E} [(\beta^n \Gamma_t^n)^2 | s_t] \leq C$  for some constant  $C$  and every  $t, n$ , then after applying COP-TD(0,  $\beta$ ),  $\hat{\rho}_{d,t}$  converges to  $\rho_d$  almost surely, and  $\theta_t$  converges to the fixed point of  $\Pi_\pi T_\pi V$ .*

Notice that COP-TD(0,  $\beta$ ) in its current form is infeasible in problems with large state spaces. Like TD( $\lambda$ ), we can introduce function approximation: represent  $\rho_d(s) \approx \theta_\rho^\top \phi_\rho(s)$  where  $\theta_\rho$  is a

1. Except full IS, however its variance is too high to be applicable in practice.

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**Algorithm 1** COP-TD(0,β), Input:  $\theta_0, \hat{\rho}_{d,0}$ ,

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- 1: Init:  $F_0 = 0$ ,  $n_0^\beta = 1$ ,  $N(s) = 0$
  - 2: **for**  $t = 1, 2, \dots$  **do**
  - 3:   Observe  $s_t, a_t, r_t, s_{t+1}$   
    Update normalization terms
  - 4:    $n_t^\beta = \beta n_{t-1}^\beta + 1$ ,  $N(s_t) = N(s_{t-1}) + 1$ ,  $\forall s \in \mathcal{S} : \hat{d}_\mu(s) = \frac{N(s)}{n_t^\beta}$   
    Update  $\Gamma_t^n$ 's weighted average:
  - 5:    $F_t = \rho_{t-1}(\beta F_{t-1} + e_{s_{t-1}})$   
    Update & project by  $\rho_d$ 's TD error:
  - 6:    $\delta_t^d = \frac{F_t^\top \hat{\rho}_{d,t}}{n_t^\beta} - \hat{\rho}_{d,t}(s_t)$ ,  $\hat{\rho}_{d,t+1} = \Pi_{\Delta_{\hat{d}_\mu}}(\hat{\rho}_{d,t} + \alpha_t^d \delta_t^d e_{s_t})$   
    Off-policy TD(0):
  - 7:    $\delta_t = r_t + \theta_t^\top (\gamma \phi(s_{t+1}) - \phi(s_t))$ ,  $\theta_{t+1} = \theta_t + \alpha_t \hat{\rho}_{d,t+1}(s_t) \rho_t \delta_t \phi(s_t)$
  - 8: **end for**
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weight vector and  $\phi_\rho(s)$  is the off-policy feature vector and adjust the algorithm accordingly. For  $\hat{\rho}_d$  to still be contained in the set  $\Delta_{d_\mu}$ , we pose the requirement on the feature vectors:  $\phi_\rho(s) \in \mathbb{R}_+^k$ , and  $\sum_s d_\mu(s) \theta_\rho^\top \phi_\rho(s) = 1$  (noted as the simplex projection  $\Pi_{\Delta_{\mathbb{E}_\mu[\phi_\rho(s)]}}$ ). We provide the details in Algorithm 2, which also incorporates non-zero  $\lambda$  (similarly to ETD( $\lambda, \beta$ )).

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**Algorithm 2** COP-TD( $\lambda, \beta$ ) with Function Approximation, Input:  $\theta_0, \theta_{\rho,0}$ 


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- 1: Init:  $F_0 = \underline{0}$ ,  $n_0^\beta = 1$ ,  $N_\phi = \underline{0}$ ,  $e_0 = \underline{0}$
  - 2: **for**  $t = 1, 2, \dots$  **do**
  - 3:   Observe  $s_t, a_t, r_t, s_{t+1}$   
    Update normalization terms:
  - 4:    $n_t^\beta = \beta n_{t-1}^\beta + 1$ ,  $N_\phi = N_\phi + \phi_\rho(s_t)$ ,  $\hat{d}_{\phi_\rho} = \frac{N_\phi}{n_t^\beta}$   
    Update  $\Gamma_t^n$ 's weighted average:
  - 5:    $F_t = \rho_{t-1}(\beta F_{t-1} + \phi_\rho(s_{t-1}))$   
    Update & project by  $\rho_d$ 's TD error:
  - 6:    $\delta_t^d = \theta_{\rho,t-1}^\top \left( \frac{F_t}{n_t^\beta} - \phi_\rho(s_t) \right)$ ,  $\theta_{\rho,t+1} = \Pi_{\Delta_{\hat{d}_{\phi_\rho}}}(\theta_{\rho,t} + \alpha_t^d \delta_t^d \phi_\rho(s_t))$   
    Off-policy TD( $\lambda$ ):
  - 7:    $M_t = \lambda + (1 - \lambda) \theta_{\rho,t+1}^\top \phi_\rho(s_t)$ ,  $e_t = \rho_t (\lambda \gamma e_t + M_t \phi(s_{t+1}))$
  - 8:    $\delta_t = r_t + \theta_t^\top (\gamma \phi(s_{t+1}) - \phi(s_t))$   $\theta_{t+1} = \theta_t + \alpha_t \delta_t e_t$
  - 9: **end for**
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**Theorem 3** *If the step sizes hold  $\sum_t \alpha_t = \sum_t \alpha_t^d = \infty$ ,  $\sum_t (\alpha_t^2 + (\alpha_t^d)^2) < \infty$ ,  $\frac{\alpha_t}{\alpha_t^d} \rightarrow 0$ ,  $t \alpha_t^d \rightarrow 0$ , and  $\mathbb{E} [(\beta^n \Gamma_t^n)^2 | s_t] \leq C$  for some constant  $C$  and every  $t, n$ , then after applying COP-TD(0, β) with function approximation satisfying  $\phi(s) \in \mathbb{R}_+^k$ ,  $\hat{\rho}_{d,t}$  converges to the fixed point of  $\Pi_{\Delta_{\mathbb{E}_\mu[\phi_\rho(s)]}} \Pi_{\phi_\rho} Y^\beta$  denoted by  $\rho_d^{COP}$  almost surely, and  $\theta_t$  converges to the fixed point of  $\Pi_{d_\mu \circ \rho_d^{COP}} T_\pi V$ .*

A possible criticism on COP-TD(0,β) is that it is not actually consistent, since in order to be consistent the original state space needs to be small, in which case every off-policy algorithm is

consistent as well. Still, the dependence on another set of features allows to trade-off accuracy with computational power in estimating  $\rho_d$  and subsequently  $V$ .

## 5. The Logarithm Approach for Handling Long Products

Konidaris et al. (2011) suggested a statistical interpretation of TD( $\lambda$ ). They show that under several assumptions the TD( $\lambda$ ) estimate  $R_{s_t}^\lambda$  is the maximum likelihood estimator of  $V(s_t)$  given  $R_{s_t}^n$ : (1) Each  $R_{s_t}^n$  is an unbiased estimator of  $V(s_t)$  (2) The random variables  $R_{s_t}^n$  are independent and specifically uncorrelated (3) The random variables  $R_{s_t}^n$  are jointly normally distributed and (4) The variance of each  $R_{s_t}^n$  is proportional to  $\lambda^n$ .

Assumption 3 is that the sampled estimators ( $R^{(n)}, \Gamma_t^n$ ) are normally distributed. For on policy TD( $\lambda$ ), this assumption might seem not too harsh as the estimators  $R^{(n)}$  represent growing **sums** of random variables. However, in our case the estimators  $\Gamma_t^n$  are growing **products** of random variables. To correct this issue we can define new estimators using a logarithm on each  $\tilde{\Gamma}_t^n$ :

$$\log[\rho_d(s_t)] = \log \left[ \mathbb{E} \left[ \hat{\rho}_d(s_{t-m}) \prod_{k=t-m}^{t-1} \rho_k \mid s_t \right] \right] \approx \log[\hat{\rho}_d(s_{t-m})] + \sum_{k=t-m}^{t-1} \mathbb{E}[\log[\rho_k] \mid s_t]. \quad (4)$$

This approximation is crude – hence, we can relate to this method mainly as a well-motivated heuristic. Notice that this formulation resembles the standard MDP formulation, only with the corresponding "reward" terms  $\log[\rho_t]$  going backward instead of forward, and no discount factor. Unfortunately, without a discount factor we cannot expect the estimated value to converge, so we propose using an artificial one  $\gamma_{\log}$ . We can incorporate function approximation for this formulation as well. Unlike COP-TD( $\lambda, \beta$ ), we can choose the features and weights as we wish with no restriction, besides the linear constraint on the resulting  $\rho_d$  through the weight vector  $\theta_\rho$ . This can be approximately enforced by normalizing  $\theta_\rho$  using  $\frac{X}{t} \doteq \frac{1}{t} \sum_t \exp(\theta_{\rho,t}^\top \phi(s_t))$  (which should equal 1 if we were exactly correct). We call the resulting algorithm Log-COP-TD( $\lambda, \beta$ ).

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**Algorithm 3** Log-COP-TD( $\lambda, \beta$ ) with Function Approximation, Input:  $\theta_0, \theta_{\rho,0}$

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- 1: Init:  $F_0 = 0, \quad n_0(\beta) = 1, \quad N(s) = 0$
  - 2: **for**  $t = 1, 2, \dots$  **do**
  - 3:   Observe  $s_t, a_t, r_t, s_{t+1}$   
       **Update normalization terms:**
  - 4:    $n_t^\beta = \beta n_{t-1}^\beta + 1, \quad N_\phi = \gamma_{\log}(\beta N_\phi + \phi_\rho(s_t)), \quad X = X + \exp(\theta_{\rho,t}^\top \phi(s_t))$   
       **Update  $\log(\Gamma_t^n)$ 's weighted average:**
  - 5:    $F_t = \beta \gamma_{\log} F_{t-1} + n_t^\beta \log[\rho(s_{t-1})]$   
       **Update & project by  $\log(\rho_d)$ 's TD error:**
  - 6:    $\delta_t^d = \frac{F_t}{n_t^\beta} + \theta_{\rho,t}^\top \left( \frac{N_\phi}{n_t^\beta} - \phi_\rho(s_t) \right), \theta_{\rho,t+1} = \theta_{\rho,t} + \alpha_t^d \delta_t^d \phi_\rho(s_t)$   
       **Off-policy TD( $\lambda$ ):**
  - 7:    $M_t = \lambda + (1 - \lambda) \exp(\theta_{\rho,t+1}^\top \phi_\rho(s_t)) / (X/t), \quad e_t = \rho_t (\lambda \gamma e_t + M_t \phi(s_{t+1}))$
  - 8:    $\delta_t = r_t + \theta_t^\top (\gamma \phi(s_{t+1}) - \phi(s_t)), \quad \theta_{t+1} = \theta_t + \alpha_t \delta_t e_t$
  - 9: **end for**
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## 6. Experiments

We have compared our algorithm to  $\text{ETD}(\lambda, \beta)$  and  $\text{GTD}(\lambda)$ . Our method of comparison estimates the value function and finds the  $d_\pi$  weighted average of the error between  $V$  and the on-policy fixed point  $\Pi_\pi TV_\pi$ :  $\|\hat{V} - \Pi_\pi TV_\pi\|_{d_\pi}^2 = \sum_s d_\pi(s) \left[ (\theta^* - \hat{\theta})^\top \phi(s) \right]^2$ , where  $\theta^*$  is the optimal  $\theta$  obtained by on-policy TD using the target policy, and we apply  $\log$  on the output to emphasize the results.

We build two MDPs with 256/1024 states and 2 actions with uniformly distributed transition probabilities, where half of the states are with reward 0 and half with reward 1. The behavior / target policy chooses action 1 with probability 0.25 / 0.75 correspondingly, regardless of the state. The feature vectors encode each state using its binary 8-bit vector along with a free parameter, the same feature vector was used for estimating  $\rho_d$ .

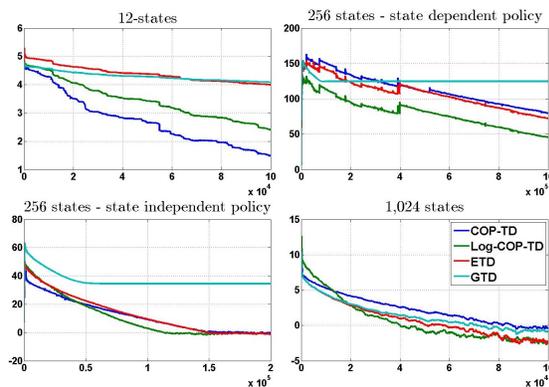
Next, we build a simple 12-states MDP designed especially to emphasize  $\text{COP-TD}(\lambda, \beta)$ 's performance and the randomized 256 states MDP with feature-dependent policy:  $\frac{\mu(a=1|s)}{\mu(a=2|s)} = \exp(\mathbf{1}^\top \phi(s)/3)$ ,  $\pi(a|s) = 1 - \mu(a|s)$ , which was designed to emphasize  $\text{Log-COP-TD}(\lambda, \beta)$ 's performance (since  $\phi(s)$  is the binary representation of  $s$ , this means  $\rho$  can have values as high as  $e^8$ ). For each algorithm we have manually chosen the best  $\lambda, \beta$  parameters from the set  $\{0, 0.25, 0.5, 0.75, 0.99\}$ .

The specialized example is a 12 states chain MDP where the agent can move either right or left. In one policy the probability to move left is 0.85 and in the other it is 0.15. This kind of setting causes distinct differences in the stationary distribution since for each policy the process is much more concentrated on a different part of the chain. The rewards were set to be 0 for the left half of the chain and 1's for the other half. The feature space aggregates each 3 consecutive states to have the same 1 in one entry and 0 in the other 3 entries.

The experiments show that on the largest MDPs  $\text{COP-TD}(\lambda, \beta)$  and  $\text{Log-COP-TD}(\lambda, \beta)$  have comparable performance to  $\text{ETD}(\lambda, \beta)$  and  $\text{GTD}(\lambda)$  for the random MDPs. On the smaller problem and the 256 state dependent policy MDP, however,  $\text{COP-TD}(\lambda, \beta)$  and  $\text{Log-COP-TD}(\lambda, \beta)$  show better results due to the high discrepancy in the stationary distributions, or very high  $\rho$  values.

## 7. Conclusion

Research on off-policy evaluation has flourished in the last decade. While a plethora of algorithms were suggested so far, these does not converge to the same point achieved by on-line TD when function approximation is required. We address this issue with  $\text{COP-TD}(\lambda, \beta)$  and proved it can achieve consistency. Despite requiring a new set of features and calibrating an additional update function,  $\text{COP-TD}(\lambda, \beta)$ 's performance does not depend as much on  $\beta$  as  $\text{ETD}(\lambda, \beta)$ , and shows promising empirical results.



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